### INSA Rennes, 4GM-AROM

# Random Models of Dynamical Systems Introduction to SDE's

## Written Exam (aka DS)

January 8, 2019

#### STOCHASTIC INTEGRAL IN INTRINSIC CLOCK = BROWNIAN MOTION

Let B be a one-dimensional standard Brownian motion, with B(0) = 0, and adapted to a given filtration  $\mathcal{F}$ , and consider the stochastic process X defined by

$$X(t) = \int_0^t \phi(u) \, dB(u) \; ,$$

for any  $t \geq 0$ , where  $\phi$  belongs to  $M_{\text{loc}}^2$ , i.e.

$$A(t) = \int_0^t |\phi(u)|^2 du < \infty ,$$

almost surely, for any  $t \geq 0$ .

(i) Write the Itô formula for the process X and for the complex-valued function  $f(x) = \exp\{i \, \lambda \, x\}$  where the scalar  $\lambda$  is fixed, between the time instants s and t, with  $0 \le s \le t$ .

For any  $t \geq 0$ , define

$$\tau(t) = \inf\{s \ge 0 : \int_0^s |\phi(u)|^2 du \ge t\}$$
,

if such time exists, and  $\tau(t) = \infty$  otherwise.

(ii) For any  $t \ge 0$ , show that the random variable  $\tau(t)$  is a stopping time, and that the equivalent definition

$$\tau(t) = \inf\{s \ge 0 : \int_0^s |\phi(u)|^2 du = t\} .$$

holds, hence  $A(\tau(t)) = t$ .

Show that  $\tau(t) \uparrow \infty$  almost surely as  $t \uparrow \infty$ .

From now on, it is assumed that

$$\int_0^\infty |\phi(u)|^2 du = \infty ,$$

i.e.  $A(t) \uparrow \infty$  almost surely as  $t \uparrow \infty$ , so that  $\tau(t) < \infty$  for any  $t < \infty$ .

- (iii) Show (a simple graphic could help to prove (a) and (b)) that
  - (a) the mapping  $t \mapsto \tau(t)$  is non–decreasing and left–continuous,
  - (b) for any  $t, s \ge 0$

$${A(t) \ge s} = {\tau(s) \le t}$$
,

(c) for any nonnegative Borel measurable function f and for any t > 0

$$\int_0^t f(\tau(s)) ds = \int_0^{\tau(t)} f(s) dA(s) .$$

[Hint: just prove (c) for any function f of the form of an indicator function, defined by  $f(s) = 1_{(0 \le s \le L)}$  for any  $s \ge 0$  and for some L > 0 (the same result for an arbitrary nonnegative Borel function would follow by a monotone class argument).]

The mapping  $\tau$  is called the *intrinsic clock* (or *intrinsic time*) for the stochastic process X, and the time–changed stochastic process Z is defined by

$$Z(t) = X(\tau(t)) = \int_0^{\tau(t)} \phi(u) dB(u) ,$$

for any  $t \geq 0$ .

(iv) Using the representation obtained in question (i) and using the result obtained in question (iii-c), show that

$$\exp\{i \,\lambda \, Z(t)\} \, = \, \exp\{i \,\lambda \, Z(s)\} + i \,\lambda \, \int_{\tau(s)}^{\tau(t)} \exp\{i \,\lambda \, X(u)\} \,\phi(u) \,dB(u)$$
$$- \, \frac{1}{2} \,\lambda^2 \, \int_s^t \exp\{i \,\lambda \, Z(u)\} \,du \,\,,$$

for any  $0 \le s \le t$ .

Introduce the  $\sigma$ -algebra  $\mathcal{A}(t) = \mathcal{F}(\tau(t))$ , i.e.  $A \in \mathcal{A}(t)$  iff for any  $u \geq 0$ 

$$A \cap \{\tau(t) \le u\} \in \mathcal{F}(u)$$
.

If the stochastic process M, defined by the stochastic integral

$$M(t) = \int_0^t \exp\{i \lambda X(u)\} \phi(u) dB(u) ,$$

for any  $t \geq 0$ , would be a martingale, and if the stopping time  $\tau(t)$  would be almost surely bounded, then the optional sampling theorem would yield

$$\mathbb{E}[M(\tau(t)) - M(\tau(s)) \mid \mathcal{A}(s)] = \mathbb{E}[\int_{\tau(s)}^{\tau(t)} \exp\{i \, \lambda \, X(u)\} \, \phi(u) \, dB(u) \mid \mathcal{A}(s)] = 0 \ ,$$

for any  $0 \le s \le t$ . The purpose of the next question is to show that the same identity holds in the more general case considered here, where  $\phi$  does belong to  $M_{\rm loc}^2$  only.

(v) Show that the stopped process  $M^{\tau(t_2)}$ , defined by  $M^{\tau(t_2)}(t) = M(t \wedge \tau(t_2))$  for any  $t \geq 0$ , is a square–integrable martingale, and that

$$\mathbb{E}\left[\int_{\tau(t_1)}^{\tau(t_2)} \exp\{i \lambda X(u)\} \phi(u) dB(u) \mid \mathcal{A}(t_1)\right] = 0.$$

for any  $0 \le t_1 \le t_2$ .

[Hint: recall that for a uniformly integrable martingale, the optional sampling theorem holds for any almost surely *finite* (and not necessarily *bounded*) stopping times.]

(vi) Show that the following identity

$$\mathbb{E}\left[\exp\left\{i\lambda\left(Z(t)-Z(s)\right)\right\} \mid \mathcal{A}(s)\right] = \exp\left\{-\frac{1}{2}\lambda^2\left(t-s\right)\right\} ,$$

holds for the conditional chraracteristic function.

Conclude that the process  $Z=(Z(t)\,,\,t\geq 0)$  is a standard Brownian motion w.r.t. the filtration  $\mathcal{A}=(\mathcal{A}(t)\,,\,t\geq 0)$ .

(vii) Show that

$$\frac{\int_0^T \phi(u) dB(u)}{\int_0^T |\phi(u)|^2 du} \longrightarrow 0$$

almost surely as  $T \uparrow \infty$ .

[Hint: use the law of large numbers for Brownian motion.]

#### SEQUENTIAL MAXIMUM LIKELIHOOD ESTIMATION

Consider the following statistical model: there exist a parametric family  $(\mathbb{P}_{\theta}, \theta \in \mathbb{R})$  of probability measures and a one–dimensional stochastic process X, such that under  $\mathbb{P}_{\theta}$  it holds

$$dX(t) = \theta b(X(t)) dt + dW_{\theta}(t)$$
,

where  $W_{\theta}$  is a standard Brownian motion, and where the drift function b satisfies the global Lipschitz and linear growth conditions.

It is assumed that the maximum likelihood estimator of the parameter  $\theta$  based on the observation of  $(X(t), 0 \le t \le T)$  in the time interval [0, T] is given by the following expression

$$\widehat{\theta}(T) = \frac{\int_0^T b(X(t)) dX(t)}{\int_0^T |b(X(t))|^2 dt}.$$

Let  $\theta_0$  denote the (unknown) true value of the parameter, and let  $\mathbb{P}_0 = \mathbb{P}_{\theta_0}$  denote the corresponding probability measure.

(viii) Show that under  $\mathbb{P}_0$  the maximum likelihood estimator satisfies

$$\widehat{\theta}(T) = \theta_0 + \frac{\int_0^T b(X(t)) dW_0(t)}{\int_0^T |b(X(t))|^2 dt} ,$$

where  $W_0$  is a standard Brownian motion.

Note that this expression cannot be used in practice, since neither is  $(W_0(t), 0 \le t \le T)$  observed (available), nor is  $\theta_0$  known. The purpose of this expression is rather to analyze the behaviour of the estimator  $\hat{\theta}(T)$ , for instance its asymptotic behaviour as  $T \uparrow \infty$ .

(ix) Show that under  $\mathbb{P}_0$  the maximum likelihood estimator is strongly consistent, i.e.  $\widehat{\theta}(T) \to \theta_0$  almost surely as  $T \uparrow \infty$ .

Actually, studying the ratio of two random variables is not so easy, and it is more convenient to study the time-changed estimator

$$\overline{\theta}(H) = \widehat{\theta}(\tau(H))$$
 where  $\tau(H) = \inf\{T \ge 0 : \int_0^T |b(X(t))|^2 dt = H\}$ .

(x) Show that under  $\mathbb{P}_0$  the time-changed maximum likelihood estimator satisfies

$$\overline{\theta}(H) = \theta_0 + \frac{1}{H} \int_0^{\tau(H)} b(X(t)) dW_0(t) .$$

The benefit of considering the time–changed maximum likelihood estimator is that the denominator is now deterministic, and the problem reduces to studying a stochastic integral under its intrinsic clock.

- (xi) Using the results obtained in the first part, show that under  $\mathbb{P}_0$  the time-changed maximum likelihood estimator
  - is strongly consistent, i.e.  $\bar{\theta}(H) \to \theta_0$  almost surely as  $H \uparrow \infty$ ,
  - is unbiased (i.e. has a mean equal to the true value  $\theta_0$ ),
  - has a (nonasymptotic) variance equal to 1/H,
  - is normally distributed, with mean  $\theta_0$  and variance 1/H.